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The Impact of Energy Prices on Green Innovation

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The Impact of Energy Prices on Green Innovation[†]

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Abstract. Based on patent data and industry specific energy prices for 18 OECD countries over 30 years we investigate on an industry level the impact of energy prices on green innovation activities. Our econometric models show that energy prices and green innovation activities are positively related and that energy prices have a significantly positive impact on the ratio of green innovations to non-green innovations. More concretely, our main model shows that a 10% increase of the average energy prices of the previous five years results in a 2.7% and 4.5% increase of the number of green patents and the ratio of green patents to non-green patents, respectively. We also find that the impact of energy prices increases with an increasing lag between energy prices and innovation activities. Robustness tests confirm the main results.

Abstract (Deutsch). In der vorliegenden Studie untersuchen wir, gestützt auf Patentdaten und industriespezifischen Energiepreisen für 18 OECD-Länder, den Einfluss von Energiepreisen auf die Innovationstätigkeit in grünen Technologien. Unsere ökonometrischen Schätzungen zeigen eine positive Korrelation zwischen Energiepreisen und grüner Innovationstätigkeit; Energiepreise haben einen signifikant positive Einfluss auf die Quote von grünen im Vergleich zu nicht-grünen Innovationen. So zeigt unser Hauptmodell, dass eine Erhöhung um 10% der durchschnittlichen Energiepreise über fünf Jahre zu einer um 2.7% höheren Anzahl grüner Patente, beziehungsweise einer um 4.5% höheren Quote von grünen Patenten im Vergleich zu anderen Patenten führt. Mit zunehmendem zeitlichen Abstand zwischen Energiepreisen und Innovationstätigkeit lässt sich ein gröserer Effekt der Energiepreise feststellen. Unsere zentralen Resultate werden durch diverse Robustheitstests bestätigt.

Keywords: Innovation; patents; environment; technological change; energy prices.

JEL classification: O30; O34; Q55.

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Table of Contents

Zusammenfassung auf Deutsch.....	1
1 Introduction.....	5
2 Conceptual Background and Hypotheses	8
3 Description of the Data	10
3.1 Measurement of green inventions based on patent statistics.....	10
3.2 OECD Stan data.....	14
3.3 IEA energy data	14
3.4 Combining the data.....	17
4 Empirical Test of Hypotheses.....	17
5 Estimation results	20
5.1 Robustness tests.....	23
6 Conclusions.....	26
7 Interpretations and Conclusions Relevant to Switzerland	27
References	31
Appendix A: Additional Tables	42
Appendix B: Tables and Figures for Switzerland	50

Zusammenfassung auf Deutsch

Ziel der vorliegenden Untersuchung ist, den Einfluss von Energiepreisen auf die Generierung von “grünen“ Innovationsaktivitäten empirisch zu untersuchen. Zu diesem Zweck werden Daten über Patentanmeldungen, Energiepreise und weitere ökonomische Kontrollgrößen aus 18 OECD-Ländern (inklusive aller wichtigsten Industrienationen) über einen Zeitraum von 30 Jahren zusammengestellt. Sämtliche Variablen sind nicht nur auf Länderebene, sondern innerhalb jedes Landes auch auf Branchenebene verfügbar, wobei wir zwischen zehn Branchen des Industriesektors unterscheiden können. Aufbauend auf diesen in ihrer Art erstmalig verfügbaren Daten lässt sich anhand unseres ökonometrischen Modells feststellen, dass eine Erhöhung der Energiepreise um 10% (im Mittel über eine Fünfjahresperiode) zu einer um 2.7% höheren Anmelderate bei grünen Patenten führt.

Unternehmungen, die grüne Innovationen anstreben, sehen sich mit dem Problem einer *doppelten Externalität* konfrontiert: einerseits lassen sich die Erträge einer erfolgreichen grünen Innovation nicht vollumfänglich durch den Innovator aneignen, da Wissen den Charakter eines öffentlichen Gutes hat; andererseits kommt der durch die Innovation generierte Nutzen (sauberere Luft, geringere CO₂-Emissionen) oftmals der gesamten Öffentlichkeit und nicht nur dem privaten Anwender der Innovation zugute. Das führt zu einer geringen individuellen Zahlungsbereitschaft für diese Produkte. Somit würde der private Sektor technologische Neuerungen, wenn sie ausschliesslich den Marktkräften überlassen bleiben, in einer aus gesellschaftlicher Sicht ungenügenden Rate produzieren. Um dies zu ändern sind politische Massnahmen zur Förderung solcher Innovationstätigkeiten vonnöten. Energiepreise sind ein potentielles Instrument einer solchen Förderpolitik, weshalb sich die vorliegende Studien mit deren Wirkung auf die Innovationstätigkeit befasst. Es gibt bereits empirische Studien, die sich mit dieser Fragestellung befassen, jedoch unterscheidet sich die vorliegende Arbeit davon, vor allem hinsichtlich der umfassenden Datenbasis.

Als Mass für den Output an grünen Innovationen einer Industriebranche dienen – wie bereits in Publikationen anderer Autoren – Patentstatistiken, die in Zusammenarbeit mit dem Eidgenössischen Institut für Geistiges Eigentum erarbeitet wurden. Mit Hilfe des von der OECD geschaffenen *Indicator of Environmental Technologies* konnten sämtliche Patentanmeldungen entweder als „grüne“ oder „nicht-grüne“ Erfindungen kategorisiert werden. Da bei der Patentanmeldungen zwar das Ursprungsland erhoben wird, nicht aber die Branchenzugehörigkeit der anmeldenden Unternehmung, wurde auf das Konkordanzschema von Schmoch et al. (2003) zurückgegriffen, womit sich die in der Patentstatistik erhobenen Technologiefelder den von uns verwendeten Branchen zuordnen lassen. Daten zu Energiepreisen sind ebenfalls nicht unmittelbar für verschiedene Branchen verfügbar, jedoch stellt die Internationale Energieagentur IEA sowohl Preisdaten für verschiedene Energieträger als auch Informationen über die effektive Nutzung dieser Energieträger nach Branchenkategorien bereit, und zwar individuell für sämtliche Länder. Dieser Datengrundlage ermöglicht uns, Energiepreise für einzelne Branchen zu konstruieren.

Um die Hypothese eines positiven Effekts von Energiepreisen auf die Generierung von grünen Innovationen empirisch zu testen, verwenden wir das ökonometrische Model der Wissensproduktionsfunktion (*Knowledge Production Functions*). In dieser Spezifikation dienen, nebst den klassischen Produktionsfaktoren Arbeit (Anzahl Beschäftigte) und Kapital (Investitionssumme), die bereits erwähnten Energiepreise als erklärende Variable. Zusätzlich kontrollieren wir für den Einfluss des bereits erarbeiteten technologischen Wissens, und zwar anhand der Anzahl der in der Vergangenheit angemeldeten Patente beider Kategorien (grün und nicht-grün). Der geschätzte Koeffizient von Energiepreisen auf die abhängige Variable (Patentanmeldungen) lässt sich in diesem Kontext als Elastizitätsmass interpretieren; fällt er statistisch signifikant aus, ist dies eine Bestätigung unserer Forschungshypothese. Dank unserer umfangreichen Datenbasis können wir auf Panel-Schätzverfahren zurückgreifen, welche auch in Gegenwart von nicht-beobachtbarer Heterogenität zwischen den einzelnen Branchen, oder von

länderspezifischen Schocks (etwa als Folge nationaler Politiken) unverzerrte Schätzungen liefern.

Die von uns bevorzugte Schätzvariante bestätigt den erwarteten positiven Effekt von Energiepreisen auf grüne Innovationen, und zwar mit einer Elastizität von 0.27, woraus sich der eingangs erwähnte Wert von 2.7% errechnet. Dieser Koeffizient ist statistisch signifikant bei einem Schwellenwert von 10%. Betrachtet man den Einfluss der Energiepreise auf den Quotienten von grünen zu nicht-grünen Patenten (anstelle der Zahl der grünen Patente), resultiert gar eine Elastizität von 0.45 (signifikant zu 5%). Mit einer Reihe von Robustheitstests überprüfen wir, ob diese Resultate auch Änderungen in der Spezifikation der Schätzung oder in der Stichprobe standhalten. So betrachten wir etwa verschiedene Unterkategorien grüner Innovationen gesondert, wiederholen unsere Berechnung der Energiepreise anhand von abweichenden Kombinationen der verwendeten Energieträger, vergleichen verschiedene Varianten zur Berechnung des existierenden Wissensbestandes oder schliessen Beobachtungen mit Ausreisern bei der abhängigen Variable aus den Schätzungen aus. Diese und weitere Tests erhärten die Stichhaltigkeit unserer Resultate.

Die erzielten Resultate können als Beleg dafür gewertet werden, dass Energiepreise als effektives Politikinstrument zur Förderung der Innovationstätigkeit im Umweltbereich dienen können, und sind weitgehend im Einklang mit Resultaten aus ähnlich gelagerten empirischen Studien. Ein Vergleich mit letzteren legt zudem nahe, dass es bedeutsame Unterschiede in der Intensität der Wirkung der Energiepreise auf grüne Innovation zwischen verschiedenen Branchen gibt, jedoch nur geringe Unterschiede zwischen verschiedenen Ländern. Ebenfalls im Einklang mit der existierenden Literatur steht die Erkenntnis, dass grüne Patentanmeldungen höher ausfallen, je grösser der Bestand an bereits erarbeiteten grünen Wissen (gemessen an der Zahl vergangener Patentanmeldungen) ist. Zwar übt auch der nicht-grüne Wissensbestand einen positiven Einfluss auf grüne Innovationen aus, allerdings in einem deutlich geringeren Mass. Dies ist Evidenz für eine gewisse Pfadabhängigkeit grüner Technologien.

Unsere Daten ermöglichen auch eine Gegenüberstellung der Innovationsperformance zwischen der Schweizerischen Industrie und (im Durchschnitt) jener der Vergleichsländer. Es zeigt sich, dass die Schweiz deutlich unterdurchschnittlich abschneidet, was die Quote von grünen gegenüber nicht-grünen Patentanmeldungen betrifft, obwohl die Gesamtzahl an grünen Erfindungen Schweizer Herkunft bedeutsam ist. Dieses Phänomen lässt sich nicht nur für den Industriesektor insgesamt, sondern auch im Einzelnen für neun der zehn betrachteten Industriebranchen feststellen. Die Tatsache, dass in der Schweiz viele forschungsintensive Branchen beheimatet sind, in welchen ein vergleichsweise geringes Potenzial für grüne Innovationen besteht (z.B. Pharma industrie), kann diesen Rückstand somit nur teilweise erklären. Auffallend ist auch, dass die vergleichsweise geringe Neigung der Schweiz zu grünen Innovationen erst gegen Ende der Beobachtungsperiode, also ab Mitte der 90er Jahre, zu Tage tritt – in früheren Jahren lag die Quote an grünen Innovationen in der Schweiz auf Höhe des weltweiten Durchschnitts oder sogar darüber. Als potentielle Erklärungen für den Rückstand der Schweiz gibt es (a) das *Opportunitätskosten-Argument*: aufgrund der in der Vergangenheit erfolgreich getätigten Innovationsanstrengungen hat sich die Schweiz eine Spezialisierung in Technologien verschafft, welche eher nicht dem grünen Bereich zuzuordnen sind. Ein Umschwenken von diesem Spezialisierungspfad auf „grünere“ Technologien wäre für die betroffenen Unternehmungen nicht ohne Umstellungskosten zu haben (wenngleich die bereits erarbeitete nicht-grüne Wissensbasis im Hinblick auf die Entwicklung grüner Technologien dennoch hilfreich ist, wie unsere Resultate zeigen). Und (b) dürfte auch die starke Zurückhaltung der Schweizer Politik in Bezug auf die explizite Förderung einzelner Technologiefelder – darunter insbesondere „grüne“ Technologien – eine Rolle gespielt haben, weshalb die Schweiz eine deutlich niedrigere Quote an Patentanmeldungen dieser Sorte im Vergleich zu vielen Ländern mit aktiveren Förderpolitiken (darunter etwa Dänemark und Deutschland) aufweist.

1 Introduction

Despite the fact that climate change should ideally increase the demand for green technologies, firms have still low incentives to invest in green technologies as there is a ‘double externality problem’ (see, e.g., Beise and Rennings 2005, Faber and Frenken 2009, Hall and Helmers 2011). Firstly, due to the public goods nature of knowledge (see, e.g., Geroski 1995, Popp 2011) and due to financial market imperfections green technology investment decisions are complex and often linked with financial constraints. Secondly, because the greatest benefits from green innovation are likely to be public rather than private, the customers’ willingness to pay for these innovations is low. In line with this literature a study by Soltmann et al. (2013) shows that economic performance is negatively affected by green inventions. This result indicates that – given the current level of green promotion – free market incentives alone are not sufficient to allow the green invention activities of industries to rise considerably. However, technological innovations are needed to solve environmental problems. “Without significant technological development of both existing low-carbon technologies and new ones, climate change is unlikely to be limited to anything like 2°C” (see Helm 2012, p. 213). Accordingly, a kind of intervention is required to stimulate green innovation activities.

The paper at hand focuses on energy prices as a measure for environmental policy¹ and investigates if energy prices are likely to contribute to increase the probability of green innovations. More concretely, we investigate if the effects of energy prices are different for ‘green’ innovations than for ‘other than green’ innovations.

Empirical research linking environmental policy and innovation is related to a small but increasing literature. A first group of studies uses pollution abatement control expenditures (PACE) to proxy for environmental regulation stringency. Brunnermeier and Cohen (2003) found for the US that PACE is positively related to environmental innovation. Based on a data

¹ Energy prices can be seen as a measure for environmental policy, since energy prices are mainly driven by international markets and their variations between countries are mainly due to taxes or other policy measures (see Aghion et al. 2012).

set that includes 17 countries Lanjouw and Mody (1996) also found a positive correlation between PACE and environmental innovation. However, the use of PACE as a measure for policy stringency in a cross-country study is questionable due to the heterogeneity in the definitions and sampling strategies (see Johnstone et al. 2012, p. 2161). To overcome this problem Johnstone et al. (2012) used survey data. Based on this data they again found that environmental innovation is positively affected by environmental policy stringency.

Most other studies overcome the problem of comparability by using energy prices as proxy for environmental regulation.² Most of them focus on a single industry. Aghion et al. (2012) investigated the significance of energy prices for technological change by looking at the car industry based on patent data over a long period in time (1978 - 2007). They found that higher energy prices increase the propensity of 'clean' innovation in the car industry. Moreover they stated that the price effect is stronger for firms with a great stock of 'dirty' patents. Newell et al. (1999) looked at the level of product characteristics in the air-conditioning industry and found that energy prices had an observable effect on energetic features of the products offered for sale. Lanzi and Sue Wing (2011) found a positive relationship between energy prices and innovations in renewable technologies in the energy sector of 23 countries.

Popp (2002) did not focus on a single industry but a single country. He looked for the USA at 11 different technologies including supply (e.g. solar energy, fuel cells) and demand technologies (e.g. recovery of waste heat for energy, heat pumps) and found that energy prices and the existing knowledge stock have a strong and significant positive effect on innovation.

² Following the findings in other fields (consumer behavior), it is likely that the elasticity effects of energy prices and energy taxes (policy instrument) have the same direction, although they are different in magnitudes. Li et al. (2012) showed that gasoline taxes reduce gasoline consumption stronger than a tax-exclusive price increase in gasoline. Davis and Kilian (2011) find that tax elasticity is much larger than price elasticity. Ghalwash (2007) compares environmental tax elasticities with traditional price elasticities and found that the effects depend on the type of energy good. Scott (2012) finds that consumers are twice as responsive to tax-driven price changes as to market-driven price changes and also Baranzini et al. (2009) find for Switzerland that consumers' reaction depend on the source of price variation. They found that the Swiss mineral oil tax increase in 1993 decreases gasoline demand by about 3.5% in addition to its direct impact through the price increase. This additional effect may result from the consumer's understanding that taxes are not a market driven variation or that gasoline tax changes receive a great deal of attention from the media, which could contribute to reinforce consumers' reaction (Li et al. 2012).

In all these studies it is unclear whether the results also hold for other industries and/or countries. Only a few studies are based on data for more than one country and more than one industry. Johnstone et al. (2010) analysed for five different renewable energy technologies how different policies (among others energy prices) did affect innovation on a certain technology. Verdolini and Galeotti (2011) investigated the impact of energy prices on technological innovation (12 technologies like in Popp 2002) for a panel of 17 countries and found a positive sign. However, as both studies are based on data that is either aggregated on country-level or technology-level, there is a concern that there may be other macro-economic shocks correlated with both innovation and the energy price (see Aghion et al. 2012, p.5).

In the study at hand we contribute to the existing literature in many respects. Firstly, we use energy prices as a proxy for environmental regulation. This allows us to generate an industry-level data set that covers the whole manufacturing sector (grouped into 10 industries), the most important countries for green invention (18 OECD countries that are responsible for more than 95% of all green patents and total patents worldwide) and this over a period of 30 years. Secondly, we use patent data to identify green and non-green inventions.³ Patent documents considered as covering green inventions are identified according to the OECD Indicator of Environmental Technologies (see OECD 2012) that distinguishes seven environmental areas, i.e. (a) general environmental management, (b) energy generation from renewable and non-fossil sources, (c) combustion technologies with mitigation potential, (d) technologies specific to climate change mitigation, (e) technologies with potential or indirect contribution to emission mitigation, (f) emission abatement and fuel efficiency in transportation, and (g) energy efficiency in buildings and lighting. If an invention can be assigned to one of these sub-groups (a to g), it is counted as a green invention; otherwise it is counted as a non-green invention. By using the Schmoch et al. (2003) concordance scheme we switch from the technology level to the industry level. This allows us to include control variables on the aggregation level of an industry (e.g.,

³ In this paper, patents and inventions are largely used synonymously.

capital and number of employees). Furthermore, we reduce the potential problem of an omitted variable bias by controlling for industry/country specific fixed effects. Thirdly, we calculate industry specific energy prices what allows us to include country specific time fixed-effects. Compared with previous studies on a more aggregated level (e.g. country level) there is no concern that there could be macro-economic shocks correlated with both innovation and the energy prices that bias our results (see Aghion et al. 2012, p.5).

With respect to our main variable we find that energy prices do stimulate both the intensity of green innovation as well as the propensity of green innovation. In our model, a 10% increase of the average energy prices of the previous five years results in a 2.7% and 4.5% increase of the number of green patents and the ratio of green patents to other patents, respectively. Knowledge about potential political instruments to stimulate innovation in this area is of large importance. As our study shows, energy prices may serve as such an instrument. An increase in energy prices may stimulate the building of a green knowledge stock that (a) would help to achieve a country's climate targets and (b) may serve as an important fundament to establish a cleantech market for which long-term growth is predicted.

2 Conceptual Background and Hypotheses

The idea that an increase in the relative price of a production factor will direct innovation efforts towards technologies that are less intensive in the production factor becoming more expensive can be attributed to Hicks (1932, as quoted e.g. in Binswanger et al. 1978): “A change in the relative prices of the factors of production is itself a spur to invention, and to invention of a particular kind - directed to economising the use of a factor which has become relatively expensive.”

This intuitively appealing assertion has been known as the induced innovation hypothesis. Subsequent research attempted to provide microeconomic foundations for this claim and to assess its relevance for traditional welfare economics (Binswanger et al. 1978, ch. 4). Induced innovation is generally thought to exacerbate the effects of externalities not properly taken into

account. In particular, the exploitation of fossil fuels has undesirable side effects as CO₂ emissions negatively affect global climate. Two harmful mechanisms are at work as a result of not having adequately priced these energy resources (by failing to take into consideration their negative externalities, e.g. by charging a CO₂ tax): price signals not only affect entrepreneurs' choice of input combinations, given the production techniques currently available; but they also affect their choice of which production technologies to develop for future use.

Taking the opposite perspective, it can be argued that taking into account induced innovation renders market-based policies to tackle climate change more efficient (or, more precisely, less costly). This is because such policies not only motivate profit-seeking firms to switch to less energy-demanding technologies that are available as of today, but these policies will induce firms to strengthen their efforts to develop such technologies for the future (see, e.g., Carraro and Siniscalco (1994) for a consideration of this point).

In line with the induced innovation hypothesis, Porter and van der Linde (1995) go as far as claiming that well-designed environmental regulation may bring about a net benefit to firms subject to such regulation. According to their argument, technological advances in process and product design triggered by such regulation often result not only in a decrease of harmful emissions (or of other undesirable ecological consequences), but also in new modes of production which are altogether more efficient, bringing about competitive gains that offset the initial private costs of complying with environmental policy. A controversial debate has subsequently been triggered about the general validity of their claims, which became to be known under the name of the Porter hypothesis. While we do not provide an empirical test for it in the present study, it should be noted that the Porter hypothesis implies that regulation triggers innovation. Thus, finding support for induced innovation can be regarded as a necessary but not sufficient condition for validating the claims made by Porter and van der Linde.

Subsequent theoretical research based on the Porter hypotheses, supports for what is known as the “weak” version of the Porter hypotheses, i.e. that energy prices are positively

related with green innovation. Mohr (2002) showed that environmental regulation, like higher energy prices, are encouraging firms to invest in clean technologies. Also Mohr and Saha (2008) showed that environmental taxes trigger green innovation. Schmutzler (2001) chose an owner-manager model and confirmed that environmental taxes lead to innovation activities if some restrictive conditions are fulfilled. Hence, we formulate the two following hypotheses:

H1: Energy prices are positively related to the number of ‘green’ innovations (i.e., the *intensity* to patent in green technologies).

H2: Energy prices are positively related to the number of green patents relative to other patents (i.e., the *propensity* to patent in green technologies).

Econometric estimations (see, e.g., Popp 2002 (for different technologies), Aghion et al. 2012 (for the car industry)) confirm the fact that energy prices are positively related with the green innovation activities. Van Leeuwen and Mohnen (2013) found strong evidence that energy prices are positively related with green innovation investments.⁴

3 Description of the Data

3.1 Measurement of green inventions based on patent statistics

We use patent statistics in order to measure the green innovation activities of an industry. Although patent statistics have many disadvantages in measuring innovation output (see Aghion et al. 2012), they are a rather good proxy for input because there is a strong relationship between the number of patents and R&D expenditure (see Griliches 1990). Despite the fact that not all inventions are patentable and smaller firms are more reluctant to patent than larger firms, patent counts are still the best available source of data on innovation activities as it is readily available and comparable across countries (Johnstone et al. 2010). This is especially true for green

⁴ Horbach et al. (2012) found that cost savings are an important driver for green innovations. This implicates that higher energy prices stimulate green innovation.

technological activities, since the OECD (2012) provides a definition of green technologies based on the patent classification.

For the paper at hand, patent information is gathered in cooperation with the Swiss Federal Institute of Intellectual Property (IPI). Green patents are a sub-group of patents that are selected according to the OECD Indicator of Environmental Technologies (see OECD 2012). Based on the International Patent Classification, the OECD definition distinguishes seven environmental areas, i.e. (a) general environmental management, (b) energy generation from renewable and non-fossil sources, (c) combustion technologies with mitigation potential, (d) technologies specific to climate change mitigation, (e) technologies with potential or indirect contribution to emission mitigation, (f) emission abatement and fuel efficiency in transportation, and (g) energy efficiency in buildings and lighting.

In order to identify our proxy for the green knowledge output of an industry, further specifications and clarifications have to be made:

(a) In order to assign patents to countries, the applicant's country of residence or the inventor's country of residence may be chosen. We assigned patents according to the applicant's address. Since only those inventions were selected for which at least one PCT (Patent Cooperation Treaty) application was filed, the applicant's address was generally available.⁵ Patent applications are costly. Hence, it is very plausible that countries for which patent applications were filed are also target markets of the invention. Accordingly, there should be a direct link between these countries and the expected market performance.

(b) We collected inventions (patent families) rather than single patents. The patent data stem from the EPO (European Patent Offices) World Patent Statistical database (PATSTAT).

⁵ We may also have used the inventor's address instead. However, there may be a risk of distorting the analysis, especially for smaller countries, because the inventor may not live in the country where the invention occurs. Conversely, by using the applicant's address the analysis may be biased by patent applications from multinationals for which the country of residence of the applicant possibly differs from the country where the invention occurred. In order to investigate if there are considerable differences, we took both the inventor's information and the applicant's information for Germany. In fact, we did not see any significant differences between the analysis based on the inventor's and applicant's address for that country.

Patents were grouped into patent families according to the PATSTAT procedure. This approach has the advantage that distortions caused by different national granting procedures and different application attitudes (USA: greater number of single applications for one invention compared to Europe) are mitigated.

(c) Only inventions were considered which were at minimum filed for patent protection under the Patent Cooperation Treaty (PCT). Fees for a PCT patent application are generally higher than for patent applications filed with national or regional patent authorities. Accordingly, applicants are expected to file inventions for patent protection under the PCT if they assume the invention to have enough commercial potential to compensate for the higher fees.

(d) Most of our model variables are classified by industrial sectors and not according to the IPC technology classes. Schmoch et al. (2003) developed a concordance scheme that links technology fields of the patent statistics to industry classes.⁶ On the basis of this concordance table we thus recoded our patent data into 10 manufacturing industry classes either at the NACE two-digit level for which also energy price data were available.⁷ In comparison with patent data at the firm level, aggregating patents on an industry level should reduce potential problems with patent waves within a firm.

(e) Our data set includes patent data from 18 countries (Australia, Austria, Belgium, Canada, Denmark, Finland, France, Germany, Ireland, Italy, Japan, Korea, the Netherlands, Spain, Sweden, Switzerland, the United Kingdom and the United States). These 18 countries account for more than 95% of all green inventions as well as all other inventions worldwide. The data set includes 10 industries that capture the whole manufacturing sector (chemicals; food and tobacco; machinery; basic metals; non-metallic minerals; paper, pulp and print; textile and

⁶ Lybbert and Zolas (2012), suggest new methods for constructing concordances. In comparing different concordance, they confirmed that on a relatively coarse level (e.g., 2 digit), the Schmoch et al. (2003) concordance enable a useful empirical policy analysis.

⁷ The concordance scheme is based on patent classification and also the OECD Indicator of Environmental Technologies (see OECD 2012) is based on the patent classification, hence, we can easily distinguish green from non-green patents on the industry level. This way we can identify for each industry class the total number of green and non-green patents.

leather; transport equipment; wood and wood products; non-specified industry). The patent data is available from 1975 onwards.⁸

Figure 1 shows the aggregated development of green inventions over time. In the beginning of our sample period, only a few green inventions were registered. The number of green inventions remained very low during the following ten years. Between 1985 and 1995, the number slightly increased. The increase was, however, not disproportionately high compared with other inventions. A sharp increase in the number of green inventions can be observed since 1995. In 2009, 29'444 green inventions were protected worldwide. Due to generally low patent activity, the ratio of green inventions to other inventions was quite unstable in the beginning of our sample period. In a second stage, the ratio stabilized between 6-8%. A disproportional increase of green inventions can be observed after 2000. By 2009, the relative importance of green inventions compared to other inventions had increased to 11.6%.

Detailed descriptive statistics for our disaggregated patent data are presented in Table 1. Nearly half of all green inventions are patented in the ‘machinery’ sector (49%). Furthermore, a considerable share is patented in the two industries ‘chemicals’ (24%) and ‘transport equipment’ (16%). The industry ‘transport equipment’ (35%) is at the same time the most green-intensive industry, followed by the two industries ‘basic metals’ (14%) and ‘non-metallic minerals’ (11%).

On the country level (see Table 1) we see larger shares of non-green patents being generated by larger countries. The USA, Japan, and Germany hold 38.5%, 14.8%, and 12.8% respectively. Switzerland has a remarkable share of 3.0%, which is about three times that of Austria or Denmark, and still about twice the share of Italy. In sum Switzerland is ranked 7th in terms of the share of other patents.

Concerning the respective shares in total green patents (see column 4 in Table 1), we see a different picture. Although the USA (29.0%), Japan (21.4%), and Germany (18.0%) also show

⁸ Actually the EPO (European Patent Office) was created in 1977/78. However, patent data are already available from 1975 onwards. The reason is that we use PCT applications, which can contain patents that are filed before 1977. Hence PCT applications can be found in PATSTAT before EPO was created.

the greatest green shares, the country ranking further down the line changes. Switzerland has 2.0% of all green patents and is ranked on the 10th position, followed by Australia (1.9%), Italy (1.5%), Denmark (1.4%), and Austria (1.1%). Compared to the ranking in other patents, Switzerland shows the greatest drop and Austria or Denmark the greatest improvement in the ranking.

The last column in Table 1 shows the ratio of green patents to non-green patents. Japan (11.6%), Germany (11.3%) and Denmark (11.0%) show the highest degree of specialisation in green patent activities, followed by Canada (10.5%) and Austria (9.7%). Switzerland (5.3%) is ranked on the 17th position just before Ireland (5.2%). In sum we see from this descriptive statistics that green innovation activities show a great heterogeneity across industries and across countries.

3.2 OECD Stan data

In order to control for important industry characteristics beside their stock of knowledge we accessed the OECD STAN database (OECD 2011). We used information on labour input (total employment) and the capital stock (gross fixed capital formation, volumes at current prices) of industries relevant for our estimations.⁹

3.3 IEA energy data

To analyse the impact of energy prices on innovation, we use information on energy prices available from the International Energy Agency's (IEA) Energy Prices and Taxes Statistics (IEA 2012a) for all 18 countries that are included in our sample. The price information is available for different energy products on a country level from 1978 onwards. To get internationally comparable information, we use total end-use prices (per toe¹⁰ including taxes) for the manufacturing sector in USD (PPP). This information is available for different energy products,

⁹ For the descriptive statistics of variables from the STAN data and other model variables see also Table A.1 in the appendix.

¹⁰ Tonne of oil equivalent; unit of energy for the practical expression of energy quantities (e.g., 1 MWh = 0.086 toe).

such as electricity, light fuel oil,¹¹ natural gas and different coal products. Figure 2 shows the development of the energy prices since 1978. We see a parallel development of the energy prices of the respective sources. Several of these products show a sharp increase at the beginning of the 80s and again from 2000 onwards. While the price of light fuel oil remarkably dropped in 2008, the electricity price further increased. Electricity is the most expensive energy source at all times, followed by light fuel oil, natural gas, steam coal, and coking coal. With the exception of electricity, energy prices doubled since 2000.

At the country level (see Figure 3) we see that electricity is most expensive in Italy, followed by Japan, Korea, and Spain. Light fuel oil is most expensive in Korea followed by Italy, Spain, and Ireland. Natural gas is most expensive in Korea, followed by Sweden, Denmark, and Japan. However, the descriptive comparisons of these prices is very limited, since prices are not available for all countries for all times. Hence, the average might be biased due to the fact that energy prices are only available at later times. This is the case e.g. for Sweden in terms of natural gas.

Besides the energy prices, the IEA collects data on consumption of the different energy products (in ktoe) on the industry level. This information is available for 10 different industries of the manufacturing sector and comes from the IEA World Energy Statistics and Balances (IEA 2012b). This allows us to calculate the relative importance of a certain energy product compared with other products on the industry level. Electricity (35%) followed by other products (28%) and natural gas (23%) are the most important energy sources. Light fuel oil, steam coal and coking coal are of minor importance in the countries we looked at (see Figure 4). If we compare natural gas, light fuel oil and electricity on an industry level across all countries and all times, we

¹¹ The IEA does also collect price information for other oil products, such as motor gasoline. However, as the number of observations is very low for these variables, we could not use this price information to construct our industry specific energy price. Our energy price should nevertheless be representative, as the energy products that could be taken into account (electricity, light fuel oil, natural gas and different coal products) make up more than 70% of total energy consumption (on average over all industries and the whole time period; see Figure 4). This figure is quite impressive, as the remaining 30% do not only include motor gasoline, but also the consumption of energy products for which no price information is collected, such as energy from biogases or heat.

see that in most industries, electricity is the most important energy source (see Figure 5). Only in the non-metallic minerals industry natural gas is more important. Natural gas is also relatively important in chemicals, food and tobacco, and in the non-specified industry.

To get industry specific energy prices, we finally multiply the energy prices with the relative importance within the industry. The industry specific energy price for an industry j , in country i at time t is defined as follows:

$$Energy_price_{ijt} = \sum_{i=1}^s w_{Eijts} * \ln(Energy_price_{its})$$

where

$$w_{Eijts} = \frac{Energy_use_{ijts}}{\sum_{k=1}^{N(s)} Energy_use_{ijtk}}$$

and

$$s \in [electricity, light fuel oil, natural gas, steam coal, coking coal].$$

The information on energy consumption as well as on energy prices is available for electricity, light fuel oil, natural gas, steam coal, and coking coal. However, due to missing values for some of the price variables, the prices used in our main model are based on the three products electricity, LFO and natural gas. Besides the fact that there are fewer missing values for these three products than for the other products, these are also the three products that show the largest relative importance in our sample (see Figure 4). However, we test the sensitivity of our results to prices that are based on other baskets of energy products as well (see Table A.6).

3.4 Combining the data

As only very few patent counts could be registered in the years before 1980, we restrict the patent sample used for regression analysis to the years 1980-2009. Accordingly, the final data set includes 18 countries, 10 industry classes and a period of 30 years. This yields a data set of 5'400 observations. Because of missing values for the other model variables, the number of observations that could be used for econometric estimations is significantly lower.

4 Empirical Test of Hypotheses

As stated by Jaffe and Palmer (1997) it is very difficult to specify a theoretically satisfying structural or reduced-form innovation equation at the industry level. Hence, we follow the framework of a knowledge production function as it was formulated by Griliches (1979) and implemented in form of a modified Cobb-Douglas model by Jaffe (1986, 1989). Similar to Jaffe (1989) we look at patents as the outcome variable but we differ in two respects, first we investigate the industry level and secondly we can distinguish between different types of knowledge inputs. We formulate the following knowledge production function for an industry j , in country i at time t :¹²

$$Green_patents_{ijt} = AL_{ijt}^\alpha K_{ijt}^\beta, \quad (1)$$

where $Green_patents$ is the number of green patents (inventions), L is the labour input and K the capital-stock, A is a constant. The parameters α and β are elasticities with respect to labour and physical capital respectively. In our model we use the industries' total number of employees as a proxy for labour (L) and the gross fixed capital formation in real terms is used to proxy physical capital (K). Ideally, one would use data on the capital stock instead of capital formation. Unfortunately, this information is only available for a few countries in the STAN database. We thus use a flow variable as a proxy for physical capital. Both variables, L and K , should be positively related with innovation activity.

¹² Other functional forms, like e.g., a translog function, would require more detailed data to describe the production process (see Griliches 1979).

Expressing (3) in logarithms yields

$$\ln(\text{Green_patents}_{ijt}) = \ln(A) + \alpha \ln(L_{ijt}) + \beta \ln(K_{ijt}). \quad (2)$$

Besides the standard input factors, the current flow of green patents should also be affected by an industry's stock of knowledge. To capture this effect we augment our specification with a variable that measures an industry's stock in green patents (*Green_stock*).¹³ Following Cockburn and Griliches (1988) and Aghion et al. (2012), the patent stock is calculated using the perpetual inventory method. Following this method, the stock is defined as

$$\text{Green_stock}_{ijt} = (1 - \delta) \text{Green_stock}_{ijt-1} + \text{Green_patents}_{ijt}, \quad (3)$$

where δ is the depreciation rate of R&D capital.¹⁴ According to most of the literature, we take δ to be equal to 15% (see Keller 2002, Hall et al. 2005). However, we test the sensitivity of our results to other depreciation rates as well (see Table A.7). To capture potential effects of available knowledge in other than green technologies, we also control for the stocks of patents that are not classified as green (*Other_stock*). The stock of other patents is calculated in the same way as the stock of green patents. In line with previous literature (see, e.g., Aghion et al. 2012, Stucki and Woerter 2012) we expect that both green specific knowledge and other than green knowledge do stimulate current green innovation activities.

Finally, to test the impact of energy prices, a variable that measures the industry specific energy prices (*Energy_price*) is included in this innovation model. The augmented specification is given by:

$$\begin{aligned} \ln(\text{Green_patents}_{ijt}) = & \ln(A) + \alpha \ln(L_{ijt-1}) + \beta \ln(K_{ijt-1}) + \varphi \ln(\text{Green_stock}_{ijt-1}) \\ & + \lambda \ln(\text{Other_stock}_{ijt-1}) + \phi \ln(\text{Energy_price}_{ijt-1}) + \mu_{it} + \eta_{ij} + \varepsilon_{ijt}, \end{aligned} \quad (4)$$

¹³ Popp (2002) finds empirical evidence that failing to properly take into account measures for existing knowledge stocks may severely bias estimates of the innovation inducing effect of energy prices.

¹⁴ Due to the low number of patents before 1980, we restricted our sample period to the years 1980-2009. However, patent applications before 1980 were used to calculate the patent stocks. The initial value of the patent stock is set at $\text{Green_stock}_{1975}/(\delta+g)$, where g is the pre-1975 growth in patent stock that is assumed to be 15%.

where φ and λ are the coefficients of knowledge stocks, ϕ is the coefficient of energy prices and ε is the stochastic error term (see Table 2 for variable definition). As patent variables may contain a value of zero, we used $\ln(I+patents)$ to avoid problems with the logarithm (see Wooldridge 2002, p. 185). To deal with the potential problem of reverse causality the independent variables are introduced with a lag of one year.

To test the robustness of the price effect we use different dynamic specifications for energy prices, i.e. we use alternative lags (2-5 year lag), we construct a weighted average of past prices as proposed by Popp (2002)¹⁵ and we calculate a moving average of the energy prices of the previous five years.

To control for correlated unobserved heterogeneity, we include country specific industry fixed effects (η). This way we control for the general policy attitude of countries in terms of industry specific policy measures.¹⁶ Furthermore, to reduce the risk of an omitted variable bias from country specific shocks, we include country specific time fixed effects (μ). As stated in Aghion et al. (2012), the increase of energy prices, e.g., might be correlated with country specific subsidies for green innovation. Accordingly, the effect of energy prices may represent an indirect effect of subsidies on green innovation, and not a direct effect of prices as suggested above. The fixed effect μ captures such country specific shocks.

As we are not just interested in the effect of energy prices on the total number of green patents (i.e., the intensity of green patent activities; see H1), but also in the effect on the development of the number of green patents relative to other patents (i.e., the propensity to patent in green technologies; see H2), we alternatively estimate our innovation model with a different dependent variable that measures the difference between the logarithms of the number of green

¹⁵ As in Popp (2002), this energy price is based on an adaptive expectation model, in which expected future energy prices are a weighted average of past prices: $P_{j,t}^* = (1-\psi)\sum_{k=0}^{\infty} \psi^k P_{j,t-k}$, where ψ , the adjustment coefficient that represents the weights placed on past observations, is 0.83 (see Popp 2002 for a similar procedure), and at the beginning of the sample period where no price data for previous time periods was available, price expectations have been set to current prices.

¹⁶ More concretely, we control for policies that are industry specific and do not change across time. These fixed effects do not control for industry specific policy shocks.

patents and non-green patents (ratio of green patents to non-green patents). Our second model thus reads as follows:

$$\begin{aligned} \ln(\text{Green_patents}_{ijt}) - \ln(\text{Other_patents}_{ijt}) &= \ln(A) + \alpha \ln(L_{ijt-1}) + \beta \ln(K_{ijt-1}) \\ &+ \varphi \ln(\text{Green_stock}_{ijt-1}) + \lambda \ln(\text{Other_stock}_{ijt-1}) + \phi \ln(\text{Energy_price}_{ijt-1}) \\ &+ \mu_u + \eta_{ij} + \varepsilon_{ijt}. \end{aligned} \quad (5)$$

5 Estimation results

The main results are presented in Tables 3 and 4. Table 3 shows OLS log linear fixed-effects estimations for the number of green patents.¹⁷ The columns with uneven numbers show the results of the full model as specified in equation 4 for different dynamic specifications of the price variable. The columns with even numbers show the results for the same estimations without capital control (reduced model), which significantly increases the number of observations. To test whether this modification does lead to an omitted variable bias, Table A.2 shows the results for the reduced models based on the same observations that are available in the full model. As the results for the energy price variable do only marginally differ between these two models, we conclude that at least the result for the energy price should not be affected by an omitted variable bias in the reduced models. Table 4 shows the results for the model with the log ratio of green to non-green patents as dependent variable, as specified in equation 5.

Our econometric estimations show that energy prices are significantly positively related with green innovation. This result is in line with hypothesis H1 which states that larger energy prices stimulate current green innovation activities. Moreover, the impact of energy prices increases with an increasing time lag between energy prices and innovation activities (see Table 3).¹⁸ If we take a 5-year lag we get significant coefficients of 0.28 (full specification) and 0.27

¹⁷ Our dependent variable is the natural logarithm of the number of green patents, which is a count variable. Accordingly, count data models would be appropriate. However, these models do not allow to control for country specific time fixed effects, which is one of the main contributions of our paper. We thus decided to present as a baseline specification the OLS fixed effects regressions. Robustness tests using count data models are presented in Table A.8.

¹⁸ To test whether the differences arising from different time lags for the price variable in Table 3 are driven by the different lag structure or the different samples, Table A.3 shows the results for the same estimates, but with the same

(reduced model); if we take the moving average of energy prices over five years we see significant coefficients of 0.27 and 0.34 for the full specification and the reduced model, respectively.

Hypothesis H2 is confirmed as well, as energy prices have a significantly positive impact on the ratio of green innovation to non-green innovations (see Table 4). Accordingly, energy prices do positively affect both, the intensity and the propensity of green innovation. In our model, a 1% increase of the average energy price of the previous five years results in a 0.45% (full model) and 0.48% (reduced model) increase of the ratio of green to non-green patents. Although we do not investigate crowding out effects of green innovation activities explicitly (see van Leeuwen and Mohnen 2013, Marin 2013 for crowding out investigations on the firm-level), the relatively large difference in elasticities (0.27 in the green patents equation vs. 0.45 in the green share equation) indicates some crowding out tendencies, since the share of green innovation is more strongly affected than the number of green innovations. Indeed, we find a significantly negative effect of energy prices on non-green innovation for all but one model specification (see Table A.4).¹⁹

As described in the introduction, our model is based on a broader data set than most previous studies. It would thus be interesting to analyse how this fact does affect the impact of the energy prices. As previous models either include different control variables or even use different measures for green innovation, a direct comparison of the marginal effects of energy prices is hardly possible. Nevertheless, a comparison can provide evidence on the question whether the impact of energy prices differs substantially among countries and industries. Popp (2002) identifies an effect of energy prices on the share of green patents in total patents of 0.34 (long run elasticity) for the USA. Though we defined our share variable differently, the size of

observations across the models. As these results do only marginally differ from previous results, we conclude that differences across models are driven by dynamic effects.

¹⁹ The impact of energy prices on total patent activity is negative but only small in size and not statistically significant (these estimates are not presented here but are available on request).

the effect is quite similar to the 0.52 that we find with respect to the share of green innovation in non-green innovation when using comparable energy prices²⁰ for 13 countries (see columns 11 and 12 of Table 4). Aghion et al. (2012) analyse the effect of fuel prices on different innovation variables for the auto industry. Based on a slightly different model specification that also controls for other types of knowledge stocks, they identified elasticities of 0.97 and -0.57 for the number of ‘clean’ and ‘dirty’ patents, respectively. These elasticities are considerably larger than the figures we find for the total manufacturing sector (based on a lag structure of one year we find for the reduced model elasticities of 0.20 and -0.14, respectively). Accordingly, it seems that the dependency on energy prices in the auto industry is larger than the dependency in the other manufacturing industries. In line with our finding, they also find that the impact of energy prices increases with an increasing lag between energy prices and innovation activities. These comparisons indicate that the differences of price elasticities across countries are smaller than the respective differences across industries. However, further investigations are necessary to clarify this point.

The results for the control variables are in line with general expectations. Labour input (L) and physical capital (K) tend to be positively correlated with the number of green patents. However, we cannot observe a significant effect for these two variables with respect to the share of green patents. The propensity to patent in green technologies is neither affected by labour input nor by physical capital input. As expected a larger stock of green knowledge does stimulate current activities in green innovation. Furthermore, we find in Table 3 that knowledge in other than green technologies serves as a resource for green innovation as well – the effect of *Other_stock* on the number of green patents is significantly positive. The positive effect of green knowledge on current green innovation activities is, however, significantly larger than the positive effect of non-green knowledge. The effect of *Other_stock* on the share of green patents is significantly negative (see Table 4). Thus, it seems that due to opportunity costs, the relative

²⁰ Estimates based on a weighted average of lagged energy prices with a discount factor of 0.83.

impact of *Other_stock* on green innovation is smaller than the impact on other than green innovation.

5.1 Robustness tests

We made comprehensive tests to check the robustness of our main results presented in Tables 3 and 4. All these tests are based on the models without the capital flow variable and using moving averages of the energy prices of the previous five years (as appearing in the last column of Tables 3 and 4, respectively).

Estimates for different subcategories of green innovation

Our estimates are so far based on a quite broad definition of green inventions. Obviously, energy price shocks should, however, primarily affect inventions that are somehow related to energy reduction. To deal with this assertion, we estimate our previous model (column 14 of Table 3) separately for the seven environmental areas that are included in the OECD definition (see OECD 2012). The respective estimates are presented in Table A.5. The estimation results show that elasticities are larger for categories that we would suppose are more directly related to energy. Accordingly, the elasticity is largest for innovations in ‘technologies with potential or indirect contribution to emission mitigation’ (e.g., energy storage) and ‘energy generation from renewable and non-fossil sources’. More general green innovation such as innovation dealing with ‘technologies specified to climate change mitigation’ (e.g., CO₂ capturing) is not significantly affected by energy price shocks. Nevertheless, our overall results seem to be quite representative, as the effect of energy prices is significantly positive for all other subcategories, and does only marginally vary across the different groups (elasticities between 0.24 and 0.38 for the other six categories).

Alternative price variables

Despite the fact that our price variable includes the prices of the three most important energy products, the construction of this variable may affect the results of our estimates. To test the

robustness of our results with respect to the construction of the price variable, we alternatively estimated our main model of Tables 3 and 4 with price variables that are based on alternative baskets of energy products. As there are missing values for some product-specific energy prices, enlarging the price basket significantly reduces the number of observations that is available for the model estimation.²¹ To get comparable results for the different price baskets, we estimate all models for the same set of observations. The respective estimation results are presented in Table A.6. To be able to compare these results with previous results, columns (2) and (8) show the results for the previous estimates based on the smaller sample. The fact that the price elasticities of these estimates only marginally differ from previous estimates (0.36 vs. 0.34 for green intensity and 0.55 vs. 0.48 for green propensity) indicates that the reduction of the sample size does not significantly affect our results.

The estimates for the different price baskets show that the elasticities of our main models represent the lower limit. For all other price baskets the price elasticities are significantly larger. The largest elasticities can be observed for prices based on the three products electricity, light fuel oil and steam coal. Based on this basket we identify elasticities of 0.98 and 1.25 for the number of green patents and the ratio of green vs. non-green patents, respectively (see columns 3 and 9). The elasticities are lowest when natural gas prices are included in the basket. Due to the relatively low prices of natural gas (see Figure 3) and its relatively high weight compared with other energy products (see Figure 4), the price mixes that include natural gas tend to be lower. Accordingly, a relative increase in these prices does lead to a lower absolute increase in energy costs than an increase in other price mixes and as a consequence we observe smaller green innovation effects. Other factors that may affect the different elasticities across the different price mixes may be different factor substitutabilities. For example it may be comparatively difficult for an industry to replace electricity by another product when electricity prices increase.

²¹ While 3'448 observations are available when only the two products electricity and light fuel oil are included in the price basket, only 1'203 observations are available when we additionally include the three products natural gas, steam coal and coking coal.

Testing the robustness of the stock variables

In our main models (Tables 3 and 4) we applied a depreciation rate of 15% in order to calculate knowledge stocks. Table A.7 (columns 1 to 4) presents the results for alternative depreciation rates of 10% and 30%. The results are relatively independent of the chosen depreciation rate. The coefficients are similar and directions of the effects are identical.

Checking for outliers

Columns (5) to (8) of Table A.7 show the estimation results with regard to outliers. The distribution of inventions across industries is very heterogeneous. Consequently we run our estimation excluding the top 1% of performers and the top 5% of the performers, respectively.²² This only marginally affected our results. We thus conclude that our results are not driven by outliers.

Dealing with special characteristics of our data

To deal with the count data characteristics of the green patent flow variable, column (1) of Table A.8 shows the results for the fixed-effects Poisson model with robust standard errors as recommended by Allison and Waterman (2002) to correct for over-dispersion. Unfortunately, this procedure does not allow the inclusion of country specific time fixed effects, thus time fixed effects only have been included as the nearest best alternative specification. The estimation results with respect to energy prices are only marginally affected by this alternative estimation procedure. The effect of energy prices on green innovation remains statistically significant and positive, and the coefficient is only slightly smaller (0.20 vs. 0.34).

Column (2) of Table A.8 shows an OLS model that includes pre-sample fixed effects as proposed by Blundell et al. (1995) in order to deal with unobserved heterogeneity in the presence of lagged endogenous variables. In doing so we add the average level of patenting over the pre-

²² Our main estimates presented in Tables 3 and 4 are based on 144 groups. To check for outliers, we excluded all groups with an average clean or dirty patent stock greater THAN or equal to the top 1% and 5% of the groups, respectively. All in all, we thus dropped two and ten groups that account for 1.5% and 6.6% of the observations, respectively.

sample period 1975-1985 for both, green and other patents (both in logs), as well as two binary variables that measure whether an industry had any patent applications at all in the pre-sample period. This procedure does again slightly reduce the size of the effect of energy prices (0.15 vs. 0.34); the effect remains, however, statistically significant and positive.²³

6 Conclusions

Based on industry-level panel data, the paper at hand investigates the determinants of green patent applications of an industry. We find that energy prices do stimulate both, the intensity of green innovation as well as the propensity of green innovation. In our model, a 10% increase of the average energy prices of the previous five years results in a 2.7% and 4.5% increase of the number of green patents and the ratio of green to non-green patents, respectively. While the main focus is on the impact of energy prices, our model shows several other interesting results. Firstly, we find that available knowledge stocks serve as an innovation relevant resource for green innovation independent whether available knowledge is green specific knowledge or knowledge in non-green technologies. Secondly, as a large knowledge stock in non-green technologies represents larger opportunity costs with respect to green innovation, the effect of non-green knowledge on current green innovation is significantly smaller than the effect of green knowledge. Furthermore, the effect of non-green knowledge on the share of green patents is significantly negative.

In contrast to previous studies, our results are more general, as they are based on a broader empirical basis. While most previous studies focused on certain industries or countries, our data set includes the whole manufacturing sector and the most important countries for green innovation. Furthermore, we have reduced the problems of an omitted-variable bias by calculating industry-specific energy prices. When comparing our results with the results of

²³ We chose these models not to be our baseline specification due to the fact that the count data model does not allow to control for country time fixed effects, and the Blundell et al. (1995) procedure mainly corrects for the endogeneity of the lagged dependent variable, which is a control variable in our model.

previous studies, we found that price elasticities seem to vary primarily across industries and not across countries. Accordingly, energy prices do not seem to be equally suitable as an instrument to stimulate green innovation in different industries. Due to the limited number of observations that is available in our data set, it was unfortunately not possible to compare price elasticities across industries. However, it seems to be an interesting task for future research to identify such difference across industries in order to increase the efficiency of energy price regulations.

Despite a large future market potential, firms are probably not willing by themselves to invest in green technologies, as green innovation is still negatively related to economic performance (see Soltmann et al. 2013). Furthermore, free-riding possibilities in green technologies seem to be limited (see Stucki and Woerter 2012). Accordingly, knowledge about potential policy instruments to stimulate innovation in this area is of large importance. As our study shows, energy prices may serve as such an instrument. An increase in energy prices may stimulate the building of a green knowledge stock that (a) would help to achieve a country's climate targets and (b) may serve as an important fundament to establish a cleantech market for which long-term growth is predicted.

Future research should focus on differences of price elasticity across industries (see above) and it could also investigate in greater detail the link between investments in green technologies and investments in traditional technologies. Such crowding-out effects or the meaning of opportunity costs for a technological change would provide additional insights for policy makers. Since technological activities are global, such research would clearly benefit, if the studies could include country comparisons.

7 Interpretations and Conclusions Relevant to Switzerland

Switzerland is usually ranked among the most innovative countries in the world (see, e.g., Arvanitis et al. 2013). Our descriptive results confirm such findings when we refer to the patent activities of Swiss firms. When referring to the share of green patents in total green patents, the picture changes. Switzerland is ranked on the 10th position (see Table 1) in terms of its relative

share in total green patents and it is ranked 17th if we look at the ratio of green patents to other patents (green specialisation). This indicates that Swiss firms invest comparatively less in green technologies compared to other countries. The present section attempts to deal with the reasons for this apparent underperformance in green technologies, in particular by looking at the evolution across time and the sectoral patterns of green innovation in Switzerland, as compared to the rest of the world.

Figure B.1 shows, for Switzerland and for the time period 1977-2009, the number of green patents and the ratio of green patents to other patents. It can thus be directly compared to Figure 1, which displays the same information for all countries considered in this study. The pattern of total number of green patents is very similar between Switzerland and the rest of the world: starting from only few inventions in the late 1970s and early 1980s, the number of inventions rose roughly exponentially throughout the time period considered. The ratio of green patents to other patents in Switzerland evolved in a more erratic manner when compared to the same indicator worldwide, which can be attributed to the lower number of inventions available for the calculation of this series. However, it can be seen that, for the last fifteen years of the comparison period, the ratio of green patenting has been consistently lower in Switzerland than in the rest of the world. This is made even more evident in Figure B.2, where the Revealed Technological Advantage (RTA) of Switzerland in green technologies is shown, which is defined as the ratio of green patenting in Switzerland divided by the same ratio in the rest of the world. Values greater than one in this measure thus represent a higher degree of the home country's specialisation in the respective technology, values less than one a lower degree of specialisation. Switzerland's RTA has been fairly stable around a value of 0.6 since the late 1990s. Prior to that – in particular from 1988 to 1993 – Switzerland exhibited a similar or even higher propensity of green patenting than the rest of the world. To sum up: Switzerland has intensified research in green activities in recent years (the ratio of green to other patents rose from below 4% in the late 1990s to over 6% ten years later). Despite this recent effort, Swiss inventions have lately

remained, on average, less “green” than inventions elsewhere, due to the fact that there was a worldwide trend towards greener patenting from the turn of the century onwards.

Is it possible to gain deeper insights from looking at differences in the patenting behaviour of the ten manufacturing branches considered so far in the present study? Let us consider the three industries that account, at the world level, for nearly 90% of all green innovations: machinery, chemicals and transport equipment. These are, according to Table B.1, also the three most important green innovators within Switzerland. However, within each of these branches, the propensity to innovate green (measured by the ratio of green patents to other patents, column 5 in Table B1) in Switzerland is below the respective industry’s world average (column 6). The gap is small for machinery (6.1% as opposed to 7.3%), but quite considerable for chemicals (3.6% vs. 6.4%) and transport equipment (19.0% vs. 34.7%).²⁴ With respect to chemicals, the gap can partly be explained by the fact that this category includes pharmaceuticals, which accounts for a significant share of Swiss patenting, and which is an activity where green innovation opportunities are much smaller than in other activities belonging to the ‘chemicals’ branch (such as ‘basic chemicals’, and ‘other chemical products’). It thus seems reasonable to affirm that the composition of the Swiss manufacturing sector accounts to some extent for the lower observed propensity of Switzerland to generate green innovations, but that this composition effect alone cannot explain the entire gap.

Additional explanations are thus required. We briefly consider two of them here. Firstly, *opportunity costs*. Switzerland is very strong in non-green technologies and its companies invested comprehensively in building up the necessary knowledge base. It is very expensive and risky to reallocate funds towards new environmentally friendly technologies, even more if businesses in the field of non-green technologies are performing well. Hence, the opportunity

²⁴ With the exception of wood and wood products, all industries in Switzerland exhibit a similar pattern of a less-than-world-average propensity to patent green. Arvanitis et al. (2011) conduct a cross-country comparison of different industries’ propensities to generate green innovations, which is similar to the one presented here (but in more detail, and at a slightly more disaggregated level), and obtain similar findings.

costs of investing in environmental friendly technologies are on average greater in Switzerland than in comparable other countries. This may explain to some extent the hesitation to invest in green technologies in Switzerland. However, given the strong knowledge base in non-green technologies, and the significant positive effect of such knowledge for the generation of green innovations that we empirically have observed in the present study (see Table 3), the technological environment is favourable also for green innovations. This suggests that Swiss firms have the capabilities to react with green innovation activities upon more favourable market conditions (e.g. energy prices) for green innovations.

Secondly, *lack of policy*. Compared to other countries, the Swiss Government refrains from early policy interventions and it is also not in the tradition of public innovation promotion in Switzerland to provide direct funding for green innovation activities. This attitude is quite understandable, given the broad consensus among Swiss managers and policy makers that refraining from openly interventionist policies in the past has been a good thing. However, it is also not surprising that countries with an early policy commitment like Germany or Denmark show a better green performance, so far. From an economic perspective, what matters for policies to be effective is that they are coherent and involve a long-term commitment. Policies that affect energy prices by taking into account negative externalities (e.g. due to the carbon emissions resulting from the use of the respective energy source) are an example of a coherent policy signal that indicates the development of future green markets and that, in turn, stimulates green innovations.

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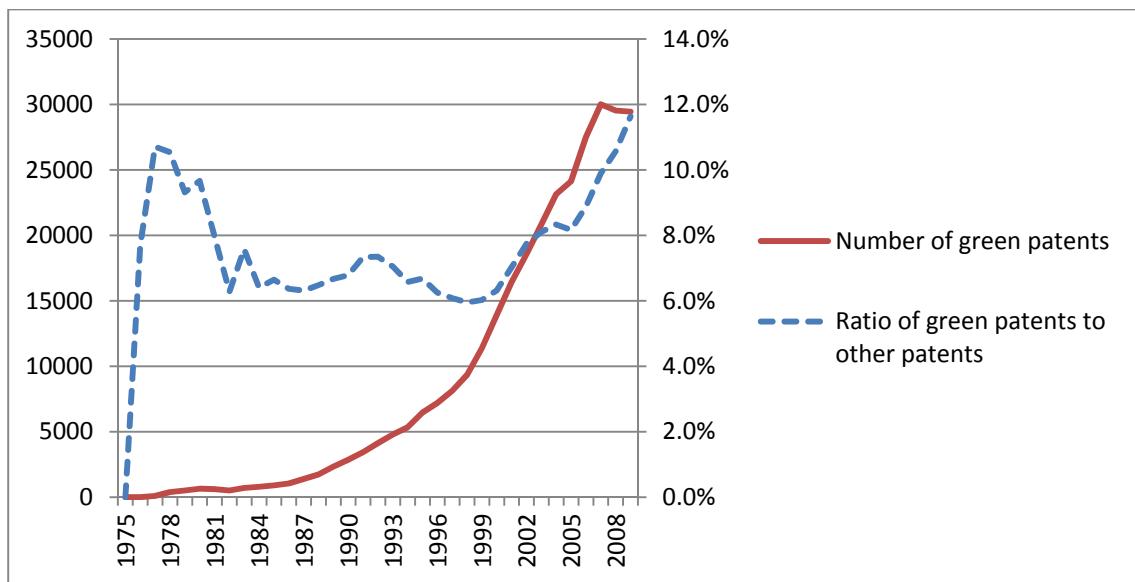
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Figure 1: Development of green patents worldwide, 1975-2009



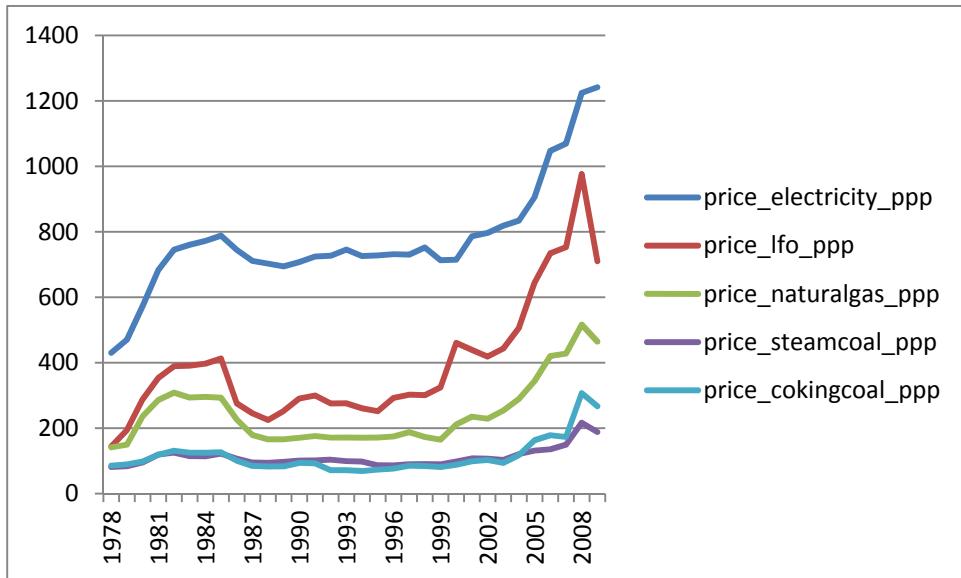
Source: Own calculations.

Table 1: Number of green and other patents (inventions) by industry and country

Period Type of patent	Other		1975-2009		Green vs. Other
	Number of other patents	Relative share in total other patents	Number of green patents	Relative share in total green patents	
Industry					
Chemicals	1174189	30.7%	75005	24.3%	6.4%
Food and tobacco	57745	1.5%	2299	0.7%	4.0%
Machinery	2057737	53.8%	151000	49.0%	7.3%
Basic metals	51937	1.4%	7058	2.3%	13.6%
Non-metallic minerals	90436	2.4%	9936	3.2%	11.0%
Paper, pulp and print	23630	0.6%	1439	0.5%	6.1%
Textile and leather	28133	0.7%	949	0.3%	3.4%
Transport equipment	145020	3.8%	50350	16.3%	34.7%
Wood and wood products	5213	0.1%	189	0.1%	3.6%
Non-specified industry	190613	5.0%	10103	3.3%	5.3%
Country					
Australia	62475	1.6%	5720	1.9%	9.2%
Austria	35787	0.9%	3479	1.1%	9.7%
Belgium	40323	1.1%	2586	0.8%	6.4%
Canada	85872	2.2%	8978	2.9%	10.5%
Switzerland	114720	3.0%	6042	2.0%	5.3%
Germany	490347	12.8%	55373	18.0%	11.3%
Denmark	38944	1.0%	4276	1.4%	11.0%
Spain	28403	0.7%	2520	0.8%	8.9%
Finland	50947	1.3%	3440	1.1%	6.8%
France	203523	5.3%	17130	5.6%	8.4%
United Kingdom	226841	5.9%	15172	4.9%	6.7%
Ireland	12425	0.3%	637	0.2%	5.1%
Italy	65926	1.7%	4640	1.5%	7.0%
Japan	565774	14.8%	65906	21.4%	11.6%
Korea	86305	2.3%	7267	2.4%	8.4%
Netherlands	130982	3.4%	8794	2.9%	6.7%
Sweden	111130	2.9%	6977	2.3%	6.3%
United States	1473929	38.5%	89391	29.0%	6.1%
Total	3824653	100%	308328	100%	8.1%

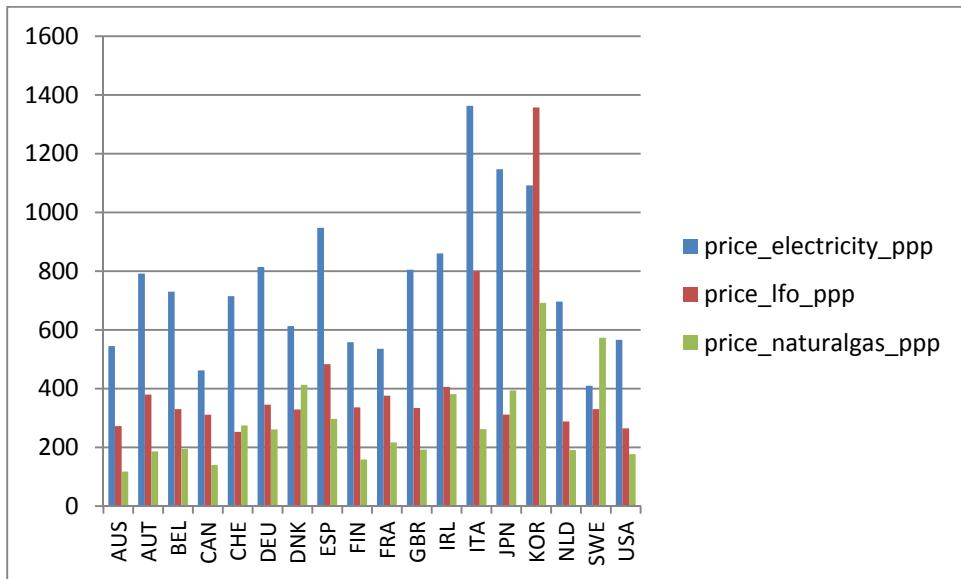
Notes: Data is based on own calculations; these statistics are based on 35 cross-sections, 18 countries and 10 industries (total of 6'300 observations); the relative share in total green patents is calculated as the share of an industry's/country's number of green patents relative to the number of all green patents in our sample (sum of green patents over all industries/countries in the sample); the ratio of green patents to other patents is defined as an industry's/ country's ratio of green patents relative to its number of other patents.

Figure 2: Energy prices for electricity, light fuel oil, natural gas, steam coal and coking coal (per tonne of oil equivalent (toe); PPP adjusted) by year, 1978-2009



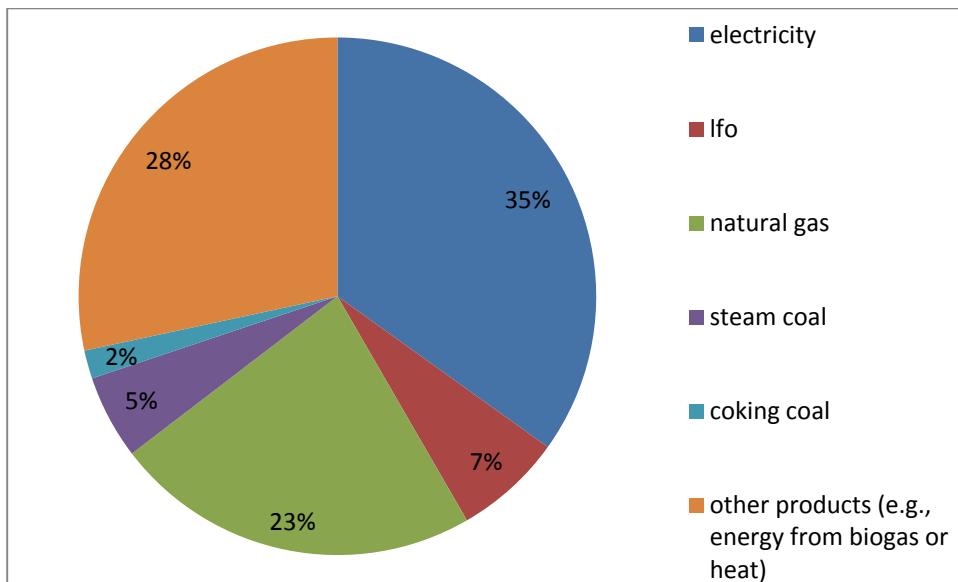
Source: IEA (2012a).

Figure 3: Average energy prices (per tonne of oil equivalent; PPP adjusted) for the three most used energy products electricity, light fuel oil and natural gas (see Figure 4) by country, 1978-2009



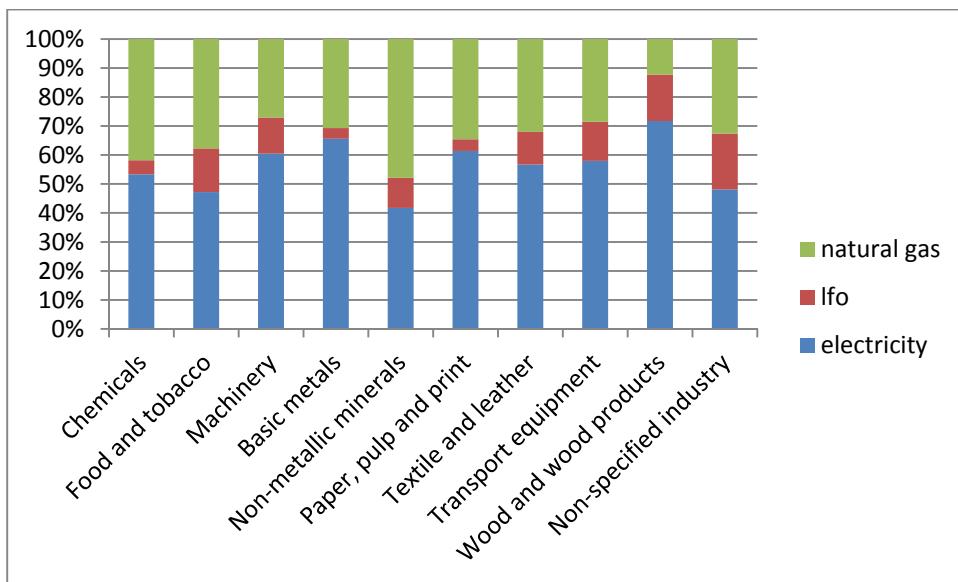
Notes: As the different price information is not available for all countries over the whole sample period, some of the figures are not directly comparable across countries and products. Natural gas prices for Sweden are for example only available for the years 2007-2009, and are thus not directly comparable with the respective prices for light fuel oil that are available for the whole sample period. Other prices averages with few observations are: Australian LFO price (6 years), Danish natural gas price (4 years) and Korean natural gas price (6 years); Source: IEA (2012a).

Figure 4: Share of total energy consumption by product, 1978-2009



Source: IEA (2012b).

Figure 5: Relative share of top three energy products by industry, 1978-2009



Source: IEA (2012b).

Table 2: Variable definition and measurement

Variable	Definition/measurement	Source
<i>Dependent variable</i>		
Green_patents _{ijt}	Number of green patents	own calculations
Other_patents _{ijt}	Number of patents that are not classified as green	own calculations
<i>Independent variable</i>		
L _{ijt}	Number of persons engaged (total employment)	OECD STAN
K _{ijt}	Gross fixed capital formation, volumes (current price value)	OECD STAN
Green_stock _{ijt}	Stock of green patents	own calculations
Other_stock _{ijt}	Stock of patents that are not classified as green	own calculations
Energy_price _{ijt}	Industry specific energy price based on electricity, light fuel oil and natural gas prices, PPP	IEA
Popp_energy_price _{ijt}	Weighted average energy prices as in Popp (2002) for the whole sample period from 1978 onwards with an adjustment coefficient of 0.83 (see Aghion et al. 2012 for a similar procedure).	IEA
Moving_average_energy_price _{ijt}	Moving average of the energy prices of the previous five years.	IEA

Table 3: Estimation results for green patent flow

Estimation method Period Dependent variable	OLS log linear fixed-effects regression													
	1981-2009							1984-2009						
	ln(Green_patents _{ijt})													
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)
ln(L _{ijt-1})	0.068 (0.080)	0.096 (0.067)	0.072 (0.078)	0.101 (0.065)	0.077 (0.083)	0.111* (0.065)	0.081 (0.082)	0.130* (0.067)	0.055 (0.079)	0.117* (0.066)	-0.030 (0.088)	0.035 (0.088)	0.028 (0.072)	0.098 (0.084)
ln(K _{ijt-1})	0.125** (0.052)		0.113** (0.052)		0.111** (0.056)		0.118** (0.058)		0.117** (0.056)		0.132** (0.054)		0.119** (0.059)	
ln(Green_stock _{ijt-1})	0.617*** (0.035)	0.613*** (0.034)	0.603*** (0.035)	0.599*** (0.035)	0.579*** (0.036)	0.580*** (0.035)	0.567*** (0.040)	0.564*** (0.037)	0.551*** (0.040)	0.552*** (0.037)	0.590*** (0.041)	0.591*** (0.039)	0.550*** (0.047)	0.551*** (0.043)
ln(Other_stock _{ijt-1})	0.150*** (0.047)	0.147*** (0.041)	0.158*** (0.047)	0.155*** (0.043)	0.139*** (0.050)	0.147*** (0.045)	0.151*** (0.057)	0.154*** (0.047)	0.180*** (0.065)	0.174*** (0.052)	0.172*** (0.051)	0.164*** (0.044)	0.158*** (0.059)	0.164*** (0.050)
ln(Energy_price _{ijt-1})	0.115 (0.089)	0.205** (0.087)												
ln(Energy_price _{ijt-2})			0.119 (0.091)	0.200** (0.087)										
ln(Energy_price _{ijt-3})				0.164* (0.091)	0.223** (0.087)									
ln(Energy_price _{ijt-4})					0.201** (0.084)	0.222*** (0.080)								
ln(Energy_price _{ijt-5})							0.277*** (0.100)	0.265*** (0.087)						
ln(Popp_energy_price _{ijt-1})									0.286 (0.174)	0.391** (0.162)				
ln(Moving_average_energy_price _{ijt-1})											0.268* (0.143)	0.342** (0.141)		
Constant	-4.475*** (1.054)	-2.812*** (0.893)	-4.483*** (1.037)	-2.920*** (0.878)	-4.657*** (1.061)	-3.179*** (0.878)	-5.073*** (1.133)	-3.436*** (0.886)	-5.307*** (1.154)	-3.540*** (0.870)	-4.296*** (1.146)	-2.984*** (1.121)	-4.985*** (1.190)	-3.880*** (1.092)
Country specific time fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes
Country specific industry fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes
N	2293	3142	2227	3051	2181	2969	2146	2899	2099	2829	1920	2725	1962	2669
Groups	126	174	126	174	116	164	116	154	116	154	105	143	116	144
R ² within	0.77	0.80	0.75	0.79	0.73	0.78	0.72	0.77	0.70	0.75	0.76	0.80	0.71	0.76

Notes: see Table 2 for the variable definitions; standard errors that are robust to heteroskedasticity and clustered at the industry-country level (clustered sandwich estimator) are in brackets under the coefficients; ***, **, * denotes statistical significance at the 1%, 5% and 10% test level, respectively.

Table 4: Estimation results for relative patenting

Estimation method Period Dependent variable	OLS log linear fixed-effects regression 1981-2009													1984-2009		
	ln(Green_patents _{ijt}) - ln(Other_patents _{ijt})															
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)		
ln(L _{ijt-1})	-0.052 (0.103)	-0.117 (0.088)	-0.045 (0.101)	-0.112 (0.088)	-0.045 (0.100)	-0.106 (0.084)	-0.039 (0.096)	-0.080 (0.081)	-0.066 (0.095)	-0.097 (0.076)	-0.158 (0.116)	-0.185* (0.095)	-0.093 (0.102)	-0.111 (0.085)		
ln(K _{ijt-1})	0.040 (0.070)	0.021 (0.069)	0.023 (0.072)	0.018 (0.070)	0.004 (0.070)	0.039 (0.068)	0.039 (0.068)	0.022 (0.068)								
ln(Green_stock _{ijt-1})	0.377*** (0.044)	0.365*** (0.042)	0.363*** (0.043)	0.350*** (0.042)	0.342*** (0.043)	0.338*** (0.044)	0.326*** (0.046)	0.320*** (0.045)	0.303*** (0.047)	0.305*** (0.045)	0.346*** (0.049)	0.333*** (0.046)	0.295*** (0.051)	0.292*** (0.051)		
ln(Other_stock _{ijt-1})	-0.313*** (0.061)	-0.368*** (0.054)	-0.302*** (0.060)	-0.356*** (0.056)	-0.280*** (0.064)	-0.337*** (0.057)	-0.227*** (0.074)	-0.300*** (0.063)	-0.231*** (0.082)	-0.294*** (0.065)	-0.301*** (0.067)	-0.359*** (0.058)	-0.227*** (0.080)	-0.297*** (0.068)		
ln(Energy_price _{ijt-1})	0.265** (0.131)	0.345*** (0.123)														
ln(Energy_price _{ijt-2})		0.253* (0.130)	0.323*** (0.122)													
ln(Energy_price _{ijt-3})			0.305** (0.134)	0.361*** (0.123)												
ln(Energy_price _{ijt-4})				0.358*** (0.131)	0.367*** (0.114)											
ln(Energy_price _{ijt-5})					0.408*** (0.142)	0.367*** (0.116)										
ln(Popp_energy_price _{ijt-1})										0.519** (0.257)	0.615** (0.238)					
ln(Moving_average_energy_price _{ijt-1})												0.450** (0.213)	0.481** (0.194)			
Constant	-3.007** (1.401)	-1.823 (1.172)	-3.007** (1.387)	-1.887 (1.169)	-3.210** (1.387)	-2.160* (1.107)	-3.709** (1.427)	-2.683** (1.093)	-3.469** (1.499)	-2.524** (1.045)	-2.986* (1.649)	-2.241 (1.524)	-3.810** (1.580)	-3.101** (1.408)		
Country specific time fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes		
Country specific industry fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes		
N	2293	3142	2227	3051	2181	2969	2146	2899	2099	2829	1920	2725	1962	2669		
Groups	126	174	126	174	116	164	116	154	116	154	105	143	116	144		
R ² within	0.50	0.50	0.48	0.48	0.47	0.47	0.43	0.44	0.43	0.43	0.51	0.51	0.43	0.44		

Notes: see Table 2 for the variable definitions; standard errors that are robust to heteroskedasticity and clustered at the industry-country level (clustered sandwich estimator) are in brackets under the coefficients; ***, **, * denotes statistical significance at the 1%, 5% and 10% test level, respectively.

Appendix A:
Additional Tables

Table A.1: Descriptive statistics

Variable	Mean	Std. Dev.	Min	Max
Green_patents _{ijt}	74.4	305.2	0	4'015
Other_patents _{ijt}	1'011.6	4'768.9	0	66'161
Energy_price _{ijt-1}	511.5	239.1	90.0	2'463.5
Green_stock _{ijt-1}	280.0	1'216.9	0	18'692.9
Other_stock _{ijt-1}	4'154.9	21'367.7	0	321'222.3
K _{ijt-1}	3.47E+09	7.09E+09	1.22E+07	8.51E+10
L _{ijt-1}	426'469.4	772'114	2'676	6'395'273

Notes: see Table 2 for the variable definitions; the set of observations is identical to that used for (1) in Table 3 (2'293 observations).

Table A.2: Identification of a possible omitted variable bias (estimates of Table 3 without capital variable but same observations)

Dependent variable	OLS log linear fixed-effects regression						
	1981-2009				1984-2009		
	ln(Green_patents _{ijt})						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
ln(L _{ijt-1})	0.150*	0.145*	0.148*	0.156*	0.123	0.054	0.103
	(0.079)	(0.076)	(0.080)	(0.083)	(0.080)	(0.087)	(0.085)
ln(Green_stock _{ijt-1})	0.622***	0.607***	0.584***	0.573***	0.557***	0.597***	0.557***
	(0.034)	(0.034)	(0.036)	(0.040)	(0.040)	(0.040)	(0.046)
ln(Other_stock _{ijt-1})	0.154***	0.163***	0.144***	0.157***	0.189***	0.179***	0.166***
	(0.047)	(0.047)	(0.051)	(0.057)	(0.066)	(0.051)	(0.059)
ln(Energy_price _{ijt-1})	0.104						
	(0.089)						
ln(Energy_price _{ijt-2})		0.112					
		(0.091)					
ln(Energy_price _{ijt-3})			0.161*				
			(0.091)				
ln(Energy_price _{ijt-4})				0.202**			
				(0.084)			
ln(Energy_price _{ijt-5})					0.283***		
					(0.099)		
ln(Popp_energy_price _{ijt-1})						0.274	
						(0.173)	
ln(Moving_average_energy_price _{ijt-1})							0.269*
							(0.141)
Constant	-2.884***	-3.013***	-3.190***	-3.604***	-3.786***	-2.519**	-3.384***
	(1.034)	(0.998)	(1.004)	(1.062)	(1.034)	(1.159)	(1.123)
Country specific time fixed effects	yes	yes	yes	yes	yes	yes	yes
Country specific industry fixed effects	yes	yes	yes	yes	yes	yes	yes
N	2293	2227	2181	2146	2099	1920	1962
Groups	126	126	116	116	116	105	116
R ² within	0.77	0.75	0.73	0.72	0.70	0.76	0.71

Notes: see Table 2 for the variable definitions; standard errors that are robust to heteroskedasticity and clustered at the industry-country level (clustered sandwich estimator) are in brackets under the coefficients; ***, **, * denotes statistical significance at the 1%, 5% and 10% test level, respectively.

Table A.3: Identification of pure dynamic effects (based on the same observations for all models)

Estimation method Period Dependent variable	OLS log linear fixed-effects regression						
	1981-2009				1984-2009		
	ln(Green_patents _{ijt})						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
ln(L _{ijt-1})	0.085 (0.075)	0.086 (0.075)	0.083 (0.075)	0.082 (0.074)	0.079 (0.074)	0.067 (0.075)	0.077 (0.074)
ln(Green_stock _{ijt-1})	0.553*** (0.046)	0.553*** (0.046)	0.552*** (0.046)	0.551*** (0.046)	0.550*** (0.046)	0.546*** (0.046)	0.548*** (0.046)
ln(Other_stock _{ijt-1})	0.174*** (0.058)	0.176*** (0.058)	0.179*** (0.057)	0.181*** (0.058)	0.185*** (0.058)	0.190*** (0.058)	0.185*** (0.058)
ln(Energy_price _{ijt-1})	0.164 (0.100)						
ln(Energy_price _{ijt-2})		0.152 (0.102)					
ln(Energy_price _{ijt-3})			0.185* (0.096)				
ln(Energy_price _{ijt-4})				0.192** (0.084)			
ln(Energy_price _{ijt-5})					0.240*** (0.089)		
ln(Popp_energy_price _{ijt-1})						0.460** (0.226)	
ln(Moving_average_energy_price _{ijt-1})							0.331** (0.153)
Constant	-2.602** (1.008)	-2.464** (0.983)	-2.698*** (0.958)	-2.674*** (0.922)	-2.879*** (0.925)	-3.672*** (1.301)	-3.562*** (1.173)
Country specific time fixed effects	yes	yes	yes	yes	yes	yes	yes
Country specific industry fixed effects	yes	yes	yes	yes	yes	yes	yes
N	2299	2299	2299	2299	2299	2299	2299
Groups	125	125	125	125	125	125	125
R ² within	0.76	0.76	0.76	0.76	0.76	0.76	0.76

Notes: see Table 2 for the variable definitions; standard errors that are robust to heteroskedasticity and clustered at the industry-country level (clustered sandwich estimator) are in brackets under the coefficients; ***, **, * denotes statistical significance at the 1%, 5% and 10% test level, respectively. Estimations are based on all observation that are available for all specifications.

Table A.4: Estimation results for other patent flow

Estimation method Period Dependent variable	OLS log linear fixed-effects regression						
	1981-2009				1984-2009		
	ln(Other_patents _{ijt})						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
ln(L _{ijt-1})	0.213*** (0.057)	0.213*** (0.062)	0.216*** (0.063)	0.210*** (0.066)	0.214*** (0.062)	0.220*** (0.064)	0.209*** (0.071)
ln(Green_stock _{ijt-1})	0.248*** (0.026)	0.248*** (0.028)	0.242*** (0.031)	0.243*** (0.033)	0.247*** (0.034)	0.258*** (0.030)	0.259*** (0.036)
ln(Other_stock _{ijt-1})	0.515*** (0.036)	0.511*** (0.039)	0.484*** (0.041)	0.454*** (0.046)	0.467*** (0.043)	0.522*** (0.038)	0.462*** (0.051)
ln(Energy_price _{ijt-1})	-0.140* (0.071)						
ln(Energy_price _{ijt-2})		-0.123* (0.067)					
ln(Energy_price _{ijt-3})			-0.138** (0.068)				
ln(Energy_price _{ijt-4})				-0.144** (0.066)			
ln(Energy_price _{ijt-5})					-0.102* (0.060)		
ln(Popp_energy_price _{ijt-1})						-0.224* (0.132)	
ln(Moving_average_energy_price _{ijt-1})							-0.139 (0.091)
Constant	-0.989 (0.701)	-1.034 (0.746)	-1.019 (0.795)	-0.753 (0.869)	-1.016 (0.821)	-0.744 (0.906)	-0.779 (0.972)
Country specific time fixed effects	yes	yes	yes	yes	yes	yes	yes
Country specific industry fixed effects	yes	yes	yes	yes	yes	yes	yes
N	3142	3051	2969	2899	2829	2725	2669
Groups	174	174	164	154	154	143	144
R ² within	0.91	0.91	0.91	0.90	0.90	0.91	0.90

Notes: see Table 2 for the variable definitions; standard errors that are robust to heteroskedasticity and clustered at the industry-country level (clustered sandwich estimator) are in brackets under the coefficients; ***, **, * denotes statistical significance at the 1%, 5% and 10% test level, respectively.

Table A.5: Estimates for different types of green innovation

Type of green patents:	OLS log linear fixed-effects regression 1984-2009 $\ln(\text{Specific_green_patents}_{ijt})$							
	General environmental management	Energy generation from renewable and non-fossil sources	Combustion technologies with mitigation potential	Technologies specific to climate change mitigation	Technologies with potential or indirect contribution to emission mitigation	Emission abatement and fuel efficiency in transportation	Energy efficiency in buildings and lighting	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	
$\ln(L_{ijt-1})$	0.053 (0.070)	0.149 (0.103)	0.036 (0.044)	-0.006 (0.039)	-0.001 (0.075)	-0.007 (0.080)	0.031 (0.061)	
$\ln(\text{Specific_green_stock}_{ijt-1})$	0.460*** (0.044)	0.526*** (0.054)	0.461*** (0.044)	0.593*** (0.050)	0.580*** (0.040)	0.547*** (0.039)	0.554*** (0.042)	
$\ln(\text{Specific_other_stock}_{ijt-1})$	0.198*** (0.052)	0.097** (0.041)	0.037 (0.029)	0.034** (0.017)	0.072* (0.039)	-0.012 (0.041)	0.078** (0.037)	
$\ln(\text{Moving_average_energy_price}_{ijt-1})$	0.239* (0.135)	0.379*** (0.133)	0.328*** (0.103)	0.074 (0.087)	0.382*** (0.115)	0.342*** (0.116)	0.255** (0.120)	
Constant	-2.851** (1.099)	-4.429*** (1.369)	-2.558*** (0.734)	-0.638 (0.699)	-2.746** (1.111)	-1.851 (1.185)	-2.442** (1.008)	
Country specific time fixed effects	yes	yes	yes	yes	yes	yes	yes	yes
Country specific industry fixed effects	yes	yes	yes	yes	yes	yes	yes	yes
N	2669	2669	2669	2669	2669	2669	2669	2669
Groups	144	144	144	144	144	144	144	144
R ² within	0.68	0.70	0.52	0.64	0.70	0.65	0.67	

Notes: see Table 2 for the variable definitions; standard errors that are robust to heteroskedasticity and clustered at the industry-country level (clustered sandwich estimator) are in brackets under the coefficients; ***, **, * denotes statistical significance at the 1%, 5% and 10% test level, respectively.

Table A.6: Estimates based on alternative price variables (same observations for all models)

Estimation method Period Dependent variable Products included in price basket	OLS log linear fixed-effects regression 1984-2009											
	ln(Green_patents _{ijt})						ln(Green_patents _{ijt}) - ln(Other_patents _{ijt})					
	electricity, light fuel oil	electricity, light fuel oil, natural gas	electricity, light fuel oil, steam coal	electricity, light fuel oil, natural gas, steam coal	electricity, light fuel oil, steam coal, coking coal	electricity, light fuel oil, natural gas, steam coal, coking coal	electricity, light fuel oil	electricity, light fuel oil, natural gas	electricity, light fuel oil, steam coal	electricity, light fuel oil, natural gas, steam coal	electricity, light fuel oil, steam coal, coking coal	electricity, light fuel oil, natural gas, steam coal, coking coal
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
ln(L _{ijt-1})	-0.059 (0.170)	-0.069 (0.157)	-0.092 (0.167)	-0.116 (0.147)	-0.080 (0.165)	-0.104 (0.146)	-0.173 (0.208)	-0.192 (0.191)	-0.212 (0.200)	-0.251 (0.181)	-0.193 (0.199)	-0.232 (0.182)
ln(Green_stock _{ijt-1})	0.437*** (0.068)	0.448*** (0.070)	0.429*** (0.069)	0.439*** (0.070)	0.431*** (0.070)	0.441*** (0.070)	0.277*** (0.087)	0.292*** (0.089)	0.270*** (0.089)	0.281*** (0.088)	0.274*** (0.090)	0.284*** (0.088)
ln(Other_stock _{ijt-1})	0.170 (0.116)	0.182 (0.115)	0.170 (0.108)	0.194* (0.114)	0.160 (0.111)	0.190 (0.115)	-0.118 (0.113)	-0.100 (0.111)	-0.119 (0.111)	-0.086 (0.105)	-0.131 (0.107)	-0.093 (0.108)
ln(Moving_average_energy_price _{ijt-1})	0.716** (0.350)	0.364** (0.182)	0.984** (0.376)	0.675** (0.274)	0.944** (0.384)	0.650** (0.267)	1.014* (0.516)	0.553* (0.284)	1.251*** (0.446)	0.929** (0.379)	1.125** (0.452)	0.871** (0.363)
Constant	-4.066 (2.681)	-1.672 (2.143)	-5.054* (2.868)	-2.825 (2.325)	-4.868* (2.890)	-2.809 (2.304)	-6.891* (3.720)	-3.793 (2.878)	-7.902** (3.143)	-5.375* (3.029)	-7.184** (3.141)	-5.175* (2.958)
Country specific time fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes
Country specific industry fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes
N	1203	1203	1203	1203	1203	1203	1203	1203	1203	1203	1203	1203
Groups	89	89	89	89	89	89	89	89	89	89	89	89
R ² within	0.80	0.80	0.80	0.80	0.80	0.80	0.31	0.31	0.32	0.32	0.31	0.32

Notes: see Table 2 for the variable definitions; standard errors that are robust to heteroskedasticity and clustered at the industry-country level (clustered sandwich estimator) are in brackets under the coefficients; ***, **, * denotes statistical significance at the 1%, 5% and 10% test level, respectively.

Table A.7: Estimates based on alternative depreciation rates and controlling for outliers, respectively

Estimation method Period Dependent variable Depreciation rate Checking for outliers	OLS log linear fixed-effects regression 1984-2009							
	ln(Green_patents _{ijt})		ln(Green_patents _{ijt}) - ln(Other_patents _{ijt})		ln(Green_patents _{ijt})		ln(Green_patents _{ijt}) - ln(Other_patents _{ijt})	
	10%	30%	10%	30%	15%	15%	15%	15%
	no	no	no	no	drop top 1%	drop top 5%	drop top 1%	drop top 5%
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
ln(L _{ijt-1})	0.101 (0.072)	0.089 (0.062)	-0.114 (0.086)	-0.103 (0.080)	0.097 (0.069)	0.091 (0.068)	-0.111 (0.085)	-0.118 (0.084)
ln(Green_stock _{ijt-1})	0.551*** (0.044)	0.539*** (0.040)	0.276*** (0.052)	0.321*** (0.047)	0.551*** (0.043)	0.548*** (0.043)	0.292*** (0.051)	0.288*** (0.051)
ln(Other_stock _{ijt-1})	0.161*** (0.054)	0.177*** (0.041)	-0.293*** (0.072)	-0.299*** (0.058)	0.163*** (0.050)	0.157*** (0.048)	-0.297*** (0.068)	-0.301*** (0.067)
ln(Moving_average_energy_price _{ijt-1})	0.356** (0.146)	0.309** (0.127)	0.491** (0.199)	0.458** (0.182)	0.352** (0.143)	0.321** (0.139)	0.486** (0.199)	0.463** (0.195)
Constant	-4.103*** (1.144)	-3.302*** (0.960)	-3.122** (1.446)	-3.176** (1.292)	-3.862*** (1.100)	-3.558*** (1.057)	-3.047** (1.432)	-2.871** (1.400)
Country specific time fixed effects	yes	yes	yes	yes	yes	yes	yes	yes
Country specific industry fixed effects	yes	yes	yes	yes	yes	yes	yes	yes
N	2669	2669	2669	2669	2629	2494	2629	2494
Groups	144	144	144	144	142	134	142	134
R ² within	0.76	0.77	0.44	0.45	0.76	0.74	0.44	0.44

Notes: see Table 2 for the variable definitions; standard errors that are robust to heteroskedasticity and clustered at the industry-country level (clustered sandwich estimator) are in brackets under the coefficients; ***, **, * denotes statistical significance at the 1%, 5% and 10% test level, respectively.

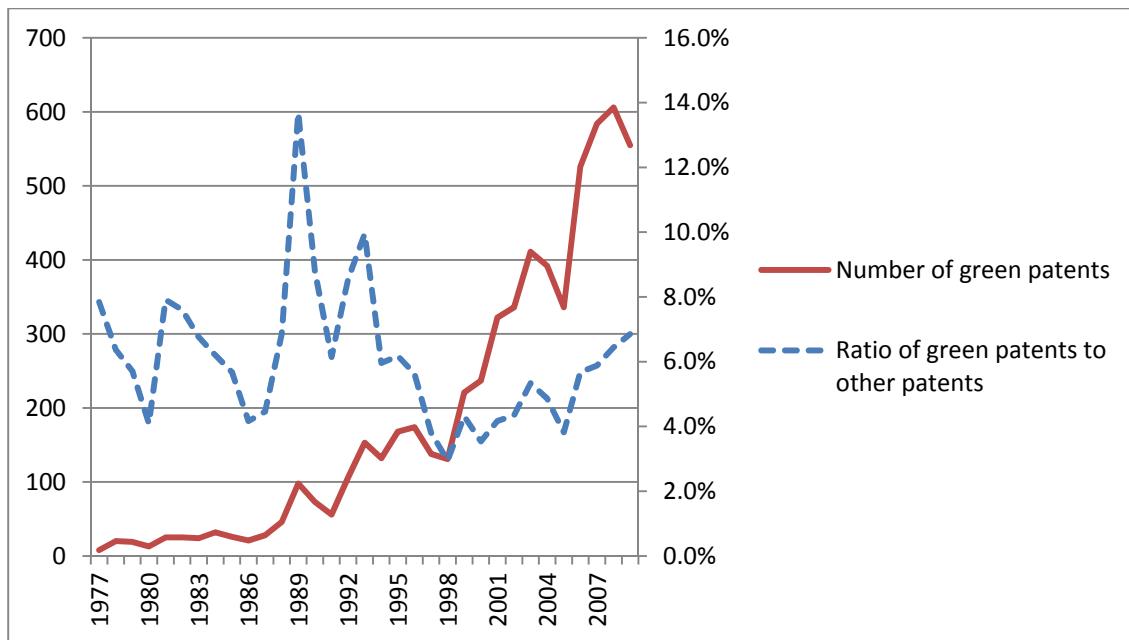
Table A.8: Models dealing with the count data characteristics of the green patent flow variable and the endogeneity of the stock variables, respectively

Estimation method	Fixed-effects Poisson regression	OLS pre-sample mean estimator
	1984-2009	
Dependent variable	Green_patents _{ijt}	ln(Green_patents _{ijt})
	(1)	(2)
ln(L _{ijt-1})	0.046 (0.121)	0.056* (0.069)
ln(Green_stock _{ijt-1})	0.798*** (0.094)	0.647*** (0.033)
ln(Other_stock _{ijt-1})	0.035 (0.119)	0.085** (0.036)
ln(Moving_average_energy_price _{ijt-1})	0.202** (0.096)	0.148** (0.069)
Constant		-2.216*** (0.715)
Year fixed effects	yes	no
Country specific industry fixed effects	yes	no
Country specific time fixed effects	no	yes
Industry fixed effects	no	yes
Pre-sample fixed effects	no	yes
N	2610	2669
Groups	137	144
Wald chi2	72782.29***	
R ²		0.94
Log Likelihood	-7674.11	

Notes: see Table 2 for the variable definitions; standard errors that are in brackets under the coefficients; ***, **, * denotes statistical significance at the 1%, 5% and 10% test level, respectively; Column (1): In line with Allison and Waterman (2002) we used robust standard errors to correct for overdispersion; Column (2): Pre-sample mean scaling approach proposed by Blundell et al. (1995) was used to account for fixed unobserved heterogeneity in the propensity to patent in the presence of lagged endogenous variables; standard errors are robust to heteroskedasticity and clustered at the industry-country level (clustered sandwich estimator).

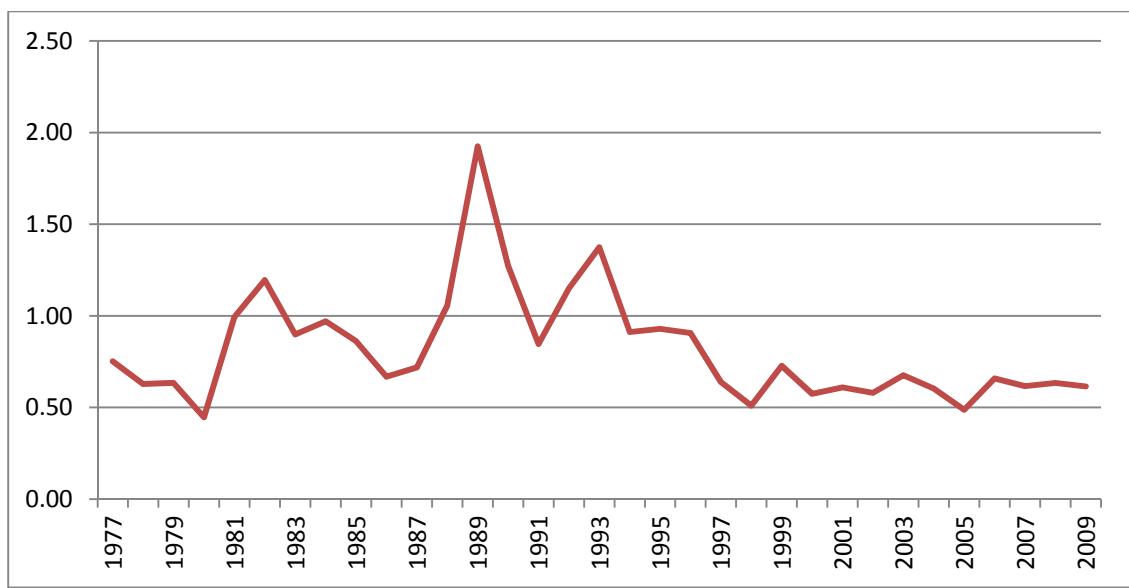
Appendix B:
Tables and Figures for Switzerland

Figure B.1: Development of green patents in Switzerland, 1977-2009



Source: Own calculations.

Figure B.2: Revealed Technological Advantage in green technology, Switzerland vs. World, 1977-2009



Source: Own calculations.

Table B.4: Number of green and other patents (inventions) by industry in Switzerland

Region Type of patent	Switzerland				World	
	Other		Green		Green vs. Other	
	Number of other patents	Relative share in total other patents	Number of green patents	Relative share in total green patents	Ratio of green patents to other patents	Ratio of green patents to other patents
Chemicals	44627	38.9%	1628	26.9%	3.6%	6.4%
Food and tobacco	3841	3.3%	132	2.2%	3.4%	4.0%
Machinery	49738	43.4%	3034	50.2%	6.1%	7.3%
Basic metals	1452	1.3%	174	2.9%	12.0%	13.6%
Non-metallic minerals	2495	2.2%	253	4.2%	10.1%	11.0%
Paper, pulp and print	1053	0.9%	36	0.6%	3.4%	6.1%
Textile and leather	1035	0.9%	21	0.3%	2.0%	3.4%
Transport equipment	2585	2.3%	491	8.1%	19.0%	34.7%
Wood and wood products	162	0.1%	7	0.1%	4.3%	3.6%
Non-specified industry	7732	6.7%	266	4.4%	3.4%	5.3%
Total	114720	100%	6042	100%	5.3%	8.1%

Notes: Data is based on own calculations; these statistics are based on 35 cross-sections and 10 industries (total of 350 observations); the relative share in total green patents is calculated as the share of an industry's number of green patents relative to the number of all green patents in our sample (sum of green patents over all industries in the sample); the ratio of green patents to other patents is defined as an industry's ratio of green patents relative to its number of other patents.